



## GALTON-WATSON BRANCHING PROBABILITY FOR THE PROCESS INEQUALITIES

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**Abstract:** The rapid development of science and technology requires the creation of mathematical models for many processes. Even the mathematical model of reducing genetic diseases transmitted from generation to generation is a branching process. This work is devoted to the application of the theory of branching of random processes to some problems .

**Key words:** branching processes, random variables, generating function.

$Z_n$ ,  $n = 0, 1, 2, \dots$  there  $p_k$  be a Galton-Watson critical process starting from a single particle,  $f(x) = \sum_{k=0}^{\infty} p_k x^k$  continuing with a single particle,  $k$  The probability of a particle.  $B(x) = f(x)$ ,  $B = B(1)$  Let (  $B$  -second-order factorial moment ) denote the radius of convergence of the series by R. The  $f(x)$  following is valid .

**Theorem 1.**  $B > 0$ ,  $R > 1$  and  $0 < y_0 < R - 1$  in that case

$$P(Z_n = k) < (1 + y_0)^{-k} \frac{1}{B_0 n/2 + 1/y_0} \quad (1)$$

here  $B_0 = B(1 + y_0)$

**Proof of Theorem 1.**  $y_n$ ,  $n = 0, 1, 2, \dots$  sequence  $0 < y_0 < R - 1$  at  $y_{n-1} = y_n + \frac{B_0}{2} y_n^2$  let the recurrence formula be satisfied.  $y_n$  the sequence decreases. Therefore

$$\frac{1}{y_n} = \frac{y_{n-1}}{y_{n-1} - y_n} = \frac{y_n + \frac{B_0}{2} y_n^2}{y_{n-1} - y_n} = \frac{1}{y_{n-1}} + \frac{B_0 y_n}{2 y_{n-1}} < \frac{1}{y_{n-1}} + \frac{B_0}{2}$$

calculating the sum

$$\frac{1}{y_n} < \frac{1}{y_0} + \frac{B_0}{2} n \quad (2)$$

we get.  $B(x)$  If we assume that the function is increasing

$$f(x) = 1 + \frac{f(1)}{1!}(x-1) + \frac{f'(1)}{2!}(x-1)^2 + \frac{f''(1)}{3!}(x-1)^3 + \dots =$$



$$= 1 + (y_n + 1 - 1) + \frac{(y_n + 1 - 1)^2}{2} f(1) + \frac{f'(1)}{3} y_n + \dots$$

$$1 + y_n + \frac{B(1 + y_n)}{2} y_n^2 < 1 + y_n + \frac{B_0}{2} y_n^2 = 1 + y_{n-1}.$$

As a result

$$f(1 + y_n) < 1 + y_{n-1},$$

$$f_2(1 + y_n) = f(f(1 + y_n)) < f(1 + y_{n-1}) < 1 + y_{n-2},$$

$$f_3(1 + y_n) < f_2(1 + y_{n-1}) < f(1 + y_{n-2}) < 1 + y_{n-3},$$

.....

$$f_n(1 + y_n) < 1 + y_0, \quad (3)$$

Here  $f_n - n$   $f$  The n interaction of and  $f_k$  the function is exponentially increasing.  $x > 1$  Let it be.  
In that case

$$P(Z_n - k) < x^{-k} f_n(x). \quad (4)$$

It is reasonable, now (4) is enough  $x = 1 + y_n$ , considering that

$$P(Z_n - k) < (1 + y_n)^{-k} f_n(1 + y_n) < (1 + y_0)^{-k} 1 + \frac{1}{B_0 n/2 + 1/y_0}$$

the theorem has been proven.

Let's look at examples of theorem 1.

**Example 1** .  $k = \sqrt{n}$ ,  $y_0 = \frac{1}{\sqrt{n}}$  In that case, the right-hand side of inequality (1) is as follows:

$$1 + \frac{1}{\sqrt{n}} - 1 + \frac{1}{\frac{B_0 n}{2} + \sqrt{n}} \sim e^{-\frac{2}{B_0 \sqrt{n}}} \rightarrow 1$$



so for large enough  $n$ ,  $P(Z_n - \sqrt{n}) < 1$  and theorem 1 will not have any useful information in this case.

Using Kolmogorov's hunting theorem

$$P(Z_n - \sqrt{n}) - P(Z_n = 0) \sim \frac{2}{nB}$$

we will have .

**Example 2.**  $y_0 = \frac{1}{n}$ ,  $k = n \ln n$  In this case, the right-hand side of equation (1) is

$$\begin{aligned} & \frac{n \ln n}{n \frac{B_0+1}{2}} \sim \frac{\frac{n \ln n}{n \frac{B_0+1}{2}}}{e^{-\frac{\ln n}{n \frac{B_0+1}{2}}}} \rightarrow 0 \\ & 1 + \frac{1}{n} - 1 + \frac{1}{\frac{B_0 n}{2} + n} \sim 1 + \frac{1}{n \frac{B_0}{2} + 1} \end{aligned}$$

So in this case,  $P(Z_n - k)$  the inequality with the exponent tends to zero.

## Literature

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